# A BRIEF STUDY ON GOMPERTZ-MAKEHAM MODEL AND SOME ASPECTS ON AGRICULTURAL GROWTH OF ASSAM 

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#### Abstract

Here we study some aspects of Gompertz-Makeham model that is used as an effective one in solving some socioeconomic development problems in Assam so to say mainly the agriculture. Fortran programme has been developed for the testing of validity of the model. Chi-Square goodness of fit test reveals what we have proposed for. It is expected that with the help of the method and the software that we have developed prediction related to growth of Agriculture of Assam and others could be made.


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## 1. INTRODUCTION

The Gompertz model is, $\mu(x)=A e^{\beta x}$ where $A$ and $\beta$ are the parameters of estimation. For our purpose we use here a redefined one, the so called Gompertz-Makeham model, $\mu(x)=A+B e^{\beta x}$ where $\mu(x)$ is the force of mortality with extra parameter A responsible for age independent component of the same. Interestingly enough this method fits well to empirical mortality distributions. It is claimed that nearly all subsequent models of the age pattern mortality have been extensions of the Gompertz-Makeham model.

The objectives of our study may be recorded as follows.
i). Fitting of the three parameters of the Gompertz-Makeham model.
ii). Estimation of the three parameters using various methods of estimations.
iii). Select the best fit model based on Chi-Square goodness of fit test.

Here we find that the method of three equidistant points used to estimate the three parameters present in the above model fits the data well compared to the method of three partial sums.

## 2. PRELIMINARIES

## PARAMETER ESTIMATION

### 2.1 DEFINITION AND METHODS

Parameter estimation is the process of calculating model parameters based on a data set. This data set can be the result of time course or steady-state experiments or both.

In the case of linear models the estimation of parameters can be done using the method of least squares. But for nonlinear models the parameters cannot be estimated using the method of least squares.

Hence we will have to go for other methods with the help of which we will be able to estimate the parameters of our desired model.
2.2 Some of the well known methods of parameter estimation are
(i) Method of three equidistant points.
(ii) Method of three partial sums.
(iii) Method of four equidistant points.
(iv) Method of four partial sums.
(v) Method of sum of reciprocals.
(vi) Composite method.

For the Gompertz Makeham model the following three methods of parameter estimation have been considered:
2.2.i Method of three equidistant points.
2.2.ii Method of three partial sums.

## 2.2. i Method of three equidistant points

In this method we determine three unknown constants $A, B$, and $\lambda$ by selecting three equidistant points such that the whole range of observation is more or less evenly covered.

Let us take three points $t_{1}, t_{2}, t_{3}$ such that

$$
\begin{aligned}
& \lambda x_{1}=t_{1}, \\
& \lambda x_{2}=t_{2}, \text { and } \\
& \lambda x_{3}=t_{3}
\end{aligned}
$$

Also let,

$$
t_{2}-t_{1}=t_{2}-t_{2}=m
$$

Let,

$$
\left\{\begin{array}{l}
\mu_{x_{1}}=A+B e^{\lambda x_{1}}=A+B e^{t_{1}}  \tag{2.2.i.1}\\
\mu_{x_{2}}=A+B e^{\lambda x_{2}}=A+B e^{t_{2}} \\
\mu_{x_{3}}=A+B e^{\lambda x_{3}}=A+B e^{t_{3}}
\end{array}\right\}
$$

Now,

$$
\begin{align*}
\mu_{x_{2}}-\mu_{x_{1}} & =B\left(e^{t_{2}}-e^{t_{1}}\right) \\
& =B\left(e^{m+t_{1}}-e^{t_{1}}\right) \\
& =B e^{t_{1}}\left(e^{m}-1\right) \tag{2.2.i.2}
\end{align*}
$$

Similarly,

$$
\begin{align*}
\mu_{x_{3}}-\mu_{x_{2}} & =B\left(e^{t_{3}}-e^{t_{2}}\right) \\
& =B\left(e^{m+t_{1}}-e^{m+t_{1}}\right) \\
& =B e^{m+t_{1}}\left(e^{m}-1\right) \tag{2.2.i.3}
\end{align*}
$$

Dividing (2.2.3) by (2.2.2),

$$
\begin{gather*}
\frac{\mu_{x_{3}}-\mu_{x_{2}}}{\mu_{x_{2}}-\mu_{x_{1}}}=e^{m} \\
\Rightarrow \frac{\log \left(\mu_{x_{3}}-\mu_{x_{2}}\right)}{\mu_{x_{2}}-\mu_{x_{1}}}=m=t_{2}-t_{1} \\
\Rightarrow \hat{\lambda}=\frac{1}{x_{2}-x_{1}} \log \left(\frac{\mu_{x_{3}}-\mu_{x_{2}}}{\mu_{x_{2}}-\mu_{x_{1}}}\right) \tag{2.2.i.4}
\end{gather*}
$$

Again,

$$
\begin{align*}
e^{m}-1 & =\frac{\mu_{x_{3}}-\mu_{x_{2}}-1}{\mu_{x_{2}}-\mu_{x_{1}}} \\
& =\frac{\mu_{x_{3}}-2 \mu_{x_{1}}+\mu_{x_{1}}}{\mu_{x_{2}}-\mu_{x_{1}}} \tag{2.2.i.5}
\end{align*}
$$

So from,(2.2.i.2) and (2.2.i.5) we have,

$$
\mu_{x_{2}}-\mu_{x_{1}}=B e^{\lambda x_{1}}\left(\frac{\mu_{x_{3}}-2 \mu_{x_{1}}+\mu_{x_{1}}}{\mu_{x_{2}}-\mu_{x_{1}}}\right)
$$

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$$
\begin{equation*}
\Rightarrow \hat{B}=\frac{\left(\mu_{x_{2}}-\mu_{x_{1}}\right)^{2}}{\mathbb{\mathbb { }}\left(\mu \rrbracket x_{3}-2 \mu_{x_{1}}+\mu_{x_{1}}\right) e^{x_{1}}}\left(\frac{\mu_{x_{3}}-\mu_{x_{1}}}{\mu_{x_{2}}-\mu_{x_{1}}}\right)^{\frac{1}{x_{2}-x_{1}}} \tag{2.2.i.6}
\end{equation*}
$$

And,

$$
\begin{equation*}
\widehat{A}=\mu_{x_{1}}-\hat{B} e^{\hat{\lambda} x_{1}} \tag{2.2.i.7}
\end{equation*}
$$

putting the value of $\widehat{\boldsymbol{B}}$ and $\widehat{\lambda}$ we get the estimated value of $\widehat{A}$. So the estimated values of the parameters $\widehat{A}, \widehat{B}, \widehat{\lambda}$ using the three equidistant points are given by (2.2.i.7), (2.2.i.6) and (2.2.i.4).

## 2.2. ii Method of three partial sums

In this method the number of observations i.e. the set of data that we will take should be multiple of three.
If N is the total number of observations then $\frac{N}{3}=m$ (say). So the three equidistant points will be:

$$
\begin{aligned}
& N_{1}=m \\
& N_{2}=N_{1}+m=2 m \\
& N_{3}=N_{2}+m=3 m
\end{aligned}
$$

Let us take,
$\mathrm{S}_{1}=$ sum of the first ' m ' observations.
$\mathrm{S}_{2}=$ sum of the second ' m ' observations.
$\mathrm{S}_{3}=$ sum of the third ' m ' observations.
Therefore,

$$
\begin{align*}
& S_{1}=m A+\frac{B e^{\lambda}\left(e^{m \lambda}-1\right)}{e^{\lambda}-1}  \tag{2.2.ii.1}\\
& S_{2}=m A+\frac{B e^{(m+1) \lambda}\left(e^{m \lambda}-1\right)}{e^{\lambda}-1}  \tag{2.2.ii.2}\\
& S_{2}=m A+\frac{B e^{(2 m+1) \lambda}\left(e^{m \lambda}-1\right)}{e^{\lambda}-1} \tag{2.2.ii.3}
\end{align*}
$$

Now,

$$
\begin{align*}
S_{2}-S_{1} & =\frac{B e^{\lambda}\left(e^{m \lambda}-1\right)^{2}}{e^{\lambda}-1}  \tag{2.2.ii.4}\\
S_{3}-S_{2} & =\frac{B e^{(m+1) \lambda}\left(e^{m \lambda}-1\right)^{2}}{e^{\lambda}-1} \tag{2.2.ii.5}
\end{align*}
$$

Dividing (2.2.ii.5) by (2.2.ii.4) we get,

$$
\begin{align*}
& \frac{S_{\mathbf{z}}-S_{\mathbf{z}}}{S_{\mathbf{z}}-S_{\mathbf{1}}}=e^{m \lambda} \\
\Rightarrow & \hat{\lambda}=\frac{1}{m} \log \frac{S_{3}-S_{2}}{S_{2}-S_{1}} \tag{2.2.ii.6}
\end{align*}
$$

And,

$$
\begin{align*}
e^{m \lambda}-1 & =\frac{S_{2}-S_{2}}{S_{2}-S_{1}}-1 \\
& =\frac{S_{2}-2 S_{2}+S_{1}}{S_{2}-S_{1}} \tag{2.2.ii.7}
\end{align*}
$$

Now putting the value of $e^{m \lambda}-\mathbf{1}$ in equation (2.2.ii.4) we have,

$$
S_{2}-S_{1}=\frac{B e^{\lambda}}{\left.\left(e^{\lambda}-1\right)\left(\frac{S_{3}-2 S_{2}+S_{1}}{S_{2}-S_{1}}\right)\right]^{2}}
$$

Therefore,

$$
\begin{equation*}
\hat{B}=\frac{\left(S_{2}-S_{1}\right)^{3}\left(e^{\lambda}-1\right)}{\left(S_{3}-2 S_{2}+S_{1}\right)^{2} e^{\lambda}} \tag{2.2.ii.8}
\end{equation*}
$$

Now putting the value of $\widehat{B}$ and $e^{m \lambda}-\mathbf{1}$ in equation (2.2.ii.1) we have,

$$
\begin{gathered}
m A=S_{1}-\frac{B e^{\lambda}\left(e^{m \lambda}-1\right)}{e^{\lambda}-1} \\
\Rightarrow m A=S_{1}-\frac{\left(S_{2}-S_{1}\right)^{2}}{S_{2}-2 S_{2}+S_{1}} \\
=\frac{S_{3} S_{1}-S_{2}^{2}}{S_{3}-2 S_{2}+S_{1}}
\end{gathered}
$$

Therefore,

$$
\begin{equation*}
\hat{A}=\frac{S_{3} S_{1}-S_{2}^{2}}{m\left(S_{3}-2 S_{2}+S_{1}\right)} \tag{2.2.ii.9}
\end{equation*}
$$

Hence, by the method of three partial sums the parameters $\widehat{A}, \widehat{B}, \widehat{\lambda}$ of the Gompertz Makeham model are given by (2.2.ii.9), (2.2.ii.8) and (2.2.ii.6).

## 3. MAIN OBSERVATION

### 3.1 GOODNESS OF FIT TEST

A very powerful test for testing the significance of the discrepancy between theory and experiment was given by Prof. Karl Pearson in 1900 and is known as "Chi-Square goodness of fit". It enables is to find whether the deviation of the experiment from theory is just by chance or is it really due to the inadequacy of the theory to fit the observed data. In this test we compare observed values with theoretical or expected values.

If $\mathrm{O}_{\mathrm{i}}(\mathrm{i}=1,2,3$, $\qquad$ $., n)$ is the observed(experimental) frequencies and $\mathrm{E}_{\mathrm{i}}(\mathrm{i}=1,2,3$, $\qquad$ .,n) is the corresponding set of expected(theoretical or hypothetical) frequencies, then Karl Pearson's Chi-square, is given by,

$$
\chi^{2}=\sum_{i=1}^{n} \frac{\left(O_{i}-E_{i}\right)^{2}}{E_{i}}
$$

with ( $\mathrm{n}-1$ ) degrees of freedom.
This is an approximation test for large values of $n$. For the validity of 'Chi square test of 'goodness of fit' between theory and experiment, the following conditions must be satisfied:
(a) The sample observations should be independent.
(b) Constraints, if any, should be linear.
(c) N , the total frequency should be reasonably large.

It may be noted that the $\chi^{2}$ test depends only on the set of observed and expected frequencies and on degrees of freedom. It does not make assumptions regarding the parent population from which the observations are taken. $\chi^{2}$ follows Chi-square distribution with ( $\mathrm{n}-1$ ) degrees of freedom. If calculated $\chi^{2}>$ tabulated value at a particular level of significance, then $\chi^{2}$ test is significant and does not fit well to the distribution at a particular level of significance. If calculated $\chi^{2}$ <tabulated value at a particular level of significance, then $\chi^{2}$ test is not significant and fits well to the distribution at a particular level of significance. Using goodness of fit test the $\chi^{2}$ values of the following table and the estimated values of the parameters using:
(a) Estimation using three equidistant points
(b) Estimation using three partial sums, are given.

### 3.2 Fitting of Gompertz- Makeham model:

TABLE 1: Agricultural production in Assam

| $\mathrm{A}=$ | -281.318 | 92.869 |
| :---: | :---: | :---: |
| $\mathrm{~B}=$ | 305.792 | -68.886 |
| $\mathrm{C}=$ | 0.020 | -0.147 |


| year | Observed data | (a) | (b) |
| :---: | :---: | :---: | :---: |
| 1988 | 30.700 | 30.700 | 33.425 |
| 1989 | 39.900 | 37.053 | 41.572 |
| 1990 | 53.000 | 43.536 | 48.603 |
| 1991 | 56.900 | 50.151 | 54.670 |
| 1992 | 66.200 | 56.900 | 59.906 |
| 1993 | 65.600 | 63.787 | 64.424 |
| 1994 | 71.900 | 70.814 | 68.323 |
| 1995 | 77.100 | 77.984 | 71.687 |
| 1996 | 77.100 | 85.300 | 74.590 |
| 1997 | 85.300 | 92.765 | 77.096 |
|  | Chi square $=$ | 5.274 | 1.981 |

TABLE 2: Net domestic product at factor cost by economic activity in agriculture

| $\mathrm{A}=$ | 94.420 | 37.879 |
| :---: | :---: | :---: |
| $\mathrm{~B}=$ | -88.853 | -32.583 |
| $\mathrm{C}=$ | -0.020 | -0.062 |


| year | Observed values | (a) | (b) |
| :---: | :---: | :---: | :---: |
| 1 | 55.670 | 55.670 | 68.163 |
| 2 | 91.090 | 81.738 | 92.356 |
| 3 | 115.620 | 103.687 | 112.574 |
| 4 | 133.600 | 122.168 | 129.469 |
| 5 | 147.350 | 137.729 | 143.589 |
| 6 | 158.210 | 150.832 | 155.389 |
| 7 | 167.000 | 161.864 | 165.250 |
| 8 | 174.260 | 171.153 | 173.490 |
| 9 | 180.360 | 178.974 | 180.377 |
| 10 | 185.560 | 185.560 | 186.132 |
| 11 | 190.030 | 191.105 | 190.941 |
| 12 | 193.940 | 195.774 | 194.941 |
| 13 | 197.360 | 199.705 | 198.319 |
| 14 | 200.400 | 203.015 | 201.126 |
| 15 | 203.110 | 205.125 | 203.472 |
| 16 | 205.540 | 208.149 | 205.432 |
| 17 | 207.730 | 210.125 | 207.071 |
| 18 | 209.720 | 211.789 | 208.440 |
| 19 | 211.530 | 213.190 | 209.584 |
| 20 | 213.190 | 214.370 | 210.540 |
|  | Chi square $=$ | 4.996 | 2.774 |

TABLE 3: General consumer price index number for agricultural labour in Assam

| $\mathrm{A}=$ | 4417.010 | 7484.693 |
| :---: | :---: | :---: |
| $\mathrm{~B}=$ | -3745.901 | -6780.670 |
| $\mathrm{C}=$ | -0.036 | -0.017 |


| Year | Observed data | (a) | (b) |
| :---: | :---: | :---: | :---: |
| 1990 | 805 | 805.000 | 821.192 |
| 1991 | 925 | 934.105 | 936.337 |
| 1992 | 1077 | 1058.596 | 1049.492 |
| 1993 | 1142 | 1178.637 | 1160.691 |
| 1994 | 1260 | 1294.387 | 1269.969 |
| 1995 | 1406 | 1406.000 | 1377.359 |
| 1996 | 1471 | 1513.623 | 1482.893 |
| 1997 | 1551 | 1617.400 | 1586.603 |
| 1998 | 1736 | 1717.468 | 1688.522 |
| 1999 | 1868 | 1813.958 | 1788.679 |
| 2000 | 1907 | 1907.000 | 1887.105 |
|  | Chi square $=$ | 8.197 | 8.109 |

TABLE 4: Food group consumer price index number for agricultural labour in Assam

| $\mathrm{A}=$ | 3245.292 | 3058.433 |
| :---: | :---: | :---: |
| $\mathrm{~B}=$ | 16.681 | 186.650 |
| $\mathrm{C}=$ | 0.401 | 0.104 |


| Year | Observed data | (a) | (b) |
| :---: | :---: | :---: | :---: |
| 1991 | 3270.200 | 3270.200 | 3265.524 |
| 1992 | 3297.200 | 3282.483 | 3288.205 |
| 1993 | 3299.700 | 3300.824 | 3313.370 |
| 1994 | 3361.100 | 3328.209 | 3341.291 |
| 1995 | 3390.100 | 3369.100 | 3372.269 |
| 1996 | 3428.200 | 3430.156 | 3406.640 |
| 1997 | 3482.200 | 3521.322 | 3444.776 |
| 1998 | 3554.800 | 3657.446 | 3487.088 |
| 1999 | 3554.800 | 3860.700 | 3534.035 |
| 2000 | 3860.700 | 4184.189 | 3586.123 |
|  | Chi square $=$ | 58.610 | 21.520 |




Table 2


Table 3


Table 4

### 3.3 CONCLUSION

From Chi square test for the set of data given in TABLE 1, TABLE 2, TABLE 3, TABLE 4 we see that the method of three equidistant points gives us the best fit than the method of three partial sums which is also clear from the graphs drawn.

### 3.5 The Fortran programming used for the estimations is given below:

```
C FITTING OF GOMPERTZ-MAKEHAM MODEL WITH THREE PARAMETERS
C Y=A+B*E**(C*T)
C 1. METHOD OF THREE EQUIDISTANT POINTS
C 2. MEHTOD OF THREE PARTIAL SUMS
    INTEGER T1,T2,T3
    DIMENSION T(40), P(40), X(40),Y(40),YCAL(40),XX(40,6)
    DIMENSION Y1(40), Y2(40), Y3(40),Z(40), PP(40)
    OPEN(UNIT=5, FILE='TEA.DAT')
    OPEN(UNIT=6, FILE='GOMPART3.DOC')
    READ(5,*) N,M
    P(I)=XX(I,1)
C WRITE(*,31) T(I),P(I)
    10 CONTINUE
    CALL THREEP(N,P,A,B,C)
    WRITE(6,55)
    WRITE(6,*) `(a): ESTIMATION USING THREE EQUIDISTANT POINTS’
    AN=N
    AA1=A
    BB1=B
    CC1=C
    DO 26 I=1,N
    Y1(I)=AA1+BB1*(EXP(CC1*I))
C
    WRITE(6,31) T(I),PP(I),Y1(I)
    CONTINUE
    CH1=0.0
    DO 301 I=1,N
    CH1=CH1+(P(I)-Y1(I))**2/ABS(Y1(I))
    CONTINUE
```

WRITE $(6,57)$ AA1,BB1,CC1,CH1
CALL PARTSUM(N,P,A,B,C)
WRITE(6,*) ‘(b): ESTIMATION USING THREE PARTIAL SUMS’
AA2=A
BB2=B
CC2=C
DO 24 I=1.N
Y2(I)=AA2+BB2*(EXP(CC2*I))
WRITE $(6,31) \mathrm{T}(), \mathrm{P}(), \mathrm{Y} 2()$
CONTINUE
CH2 $=0.0$
DO $302 \mathrm{I}=1 . \mathrm{N}$
CH2=CH2+(P(I)-Y2(I))**2/ABS(Y2(I))
CONTINUE
WRITE $(6,57)$ AA2,BB2,CC2,CH2
WRITE(6,*)
WRITE $(6,55)$
WRITE $(6,56)$
WRITE $(6,51)$ AA1,AA2
WRITE $(6,52)$ BB1,BB2
WRITE $(6,53) \mathrm{CC} 1, \mathrm{CC} 2$
WRITE $(6,54)$
DO $45 \mathrm{I}=1, \mathrm{~N}$
WRITE(6,31) T(I), P(I),Y1(I), Y2(I)
CONTIN UE
WRITE $(6,56)$
WRITE $(6,54)$ CH1,CH2
FORMAT(5X, 4F10.3)
FORMAT(20X, 'A= ‘,2F10.3)
FORMAT(20X, 'B= ',2F10.3)
FORMAT(20X, ‘C= ‘,2F10.3)
FORMAT(25X, 'Chi= ',2F10.3)
FORMAT(5X, 'FITTING OF GOMPERTZ MODEL WITH THREE PARAMETERS’/)
FORMAT (15X, '
..')
FORMAT(6X, 'A=',F10.3, 'B=’,F10.3, ‘C=',F10.3,
'Chi-Sq=’,F10.3 /)
STOP
END
SUBROUTINE THREEP(N,P,A,B,C)
C FITTING OF MODEL USING THREE EQUIDISTANT POINTS

```
C FITTING OF MODEL USING THREE PARTIAL SUMS
DIMENSION \(\mathrm{T}(40), \mathrm{P}(40), \mathrm{X}(40), \mathrm{Y}(40)\)
AN=N
```


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N1=AN/3
N2=N1+N1
$\mathrm{N} 3=\mathrm{N} 2+\mathrm{N} 1$
AAN1=N1
AN11=1/AAN1
S1=0.0
S2 $=0.0$
S3=0.0
DO 21 I=1,N
S1=S1+P(I)
CONTINUE
DO 22 I=N1+1,N2
S2=S2+P(I)
CONTINUE
C=AN11*LOG((S3-S2)/(S2-S1))
U1=((S2-S1)**3*((S3-S2)**AN11-(S2-S1)**AN11)
U2=((S3-(2*S2)+S1)**2*((S3-S2)**AN11)
B=U1/U2
C $\left.\quad \mathrm{B}=\left(\left((\mathrm{S} 2-\mathrm{S} 1)^{* *} 3\right) /(\mathrm{S} 3-(2 * \mathrm{~S} 2)+\mathrm{S} 1)^{* *} 2\right) *\left(1-\left(((\mathrm{S} 2-\mathrm{S} 1) /(\mathrm{S} 3-\mathrm{S} 2))^{* *} \mathrm{AN} 11\right)\right)\right)$
A=(((S1*S3)-(S2*S2))/(S3-(2*S2)+S1)**AN11
RETURN
END

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